The asymptotic behavior of the Hurwitz binomial distribution^{*}

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Abstract

Hurwitz's extension of Abel's binomial theorem defines a probability distribution on the set of integers from 0 to n. This is the distribution of the number of non-root vertices of a fringe subtree of a suitably defined random tree with n + 2vertices. The asymptotic behaviour of this distribution is described in a limiting regime where the distribution of the delabeled fringe subtree approaches that of a Galton-Watson tree with a mixed Poisson offspring distribution.

1 Introduction and statement of results

Hurwitz [10] discovered the following identity of polynomials in n + 2 variables x, y and $z_s, s \in [n] := \{1, \ldots, n\}$, which reduces to the binomial expansion of $(x + y)^n$ when

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 $z_s \equiv 0$:

$$\sum_{A \subseteq [n]} x (x + z_A)^{|A| - 1} (y + z_{\bar{A}})^{|\bar{A}|} = (x + y + z_{[n]})^n \tag{1}$$

where the sum is over all 2^n subsets A of [n], with the notations $z_A := \sum_{s \in A} z_s$, and |A| for the number of elements of A, and $\bar{A} := [n] - A$. See [11, 8, 19, 21] for combinatorial explanations of this identity. As observed in [15], this identity amounts to the fact that for each probability distribution p on the interval of integers $[0, n + 1] := \{0, 1, \ldots, n, n + 1\}$, the formula

$$P(V(n) = A) = p_0(p_0 + p_A)^{|A| - 1}(p_{n+1} + p_{\bar{A}})^{|\bar{A}|} \qquad (A \subseteq [n])$$
(2)

defines the probability distribution of a random subset V(n) of [n], which can be constructed as follows. As a consequence of Cayley's multinomial expansion [4, 18, 15], a probability distribution for a random tree $\mathcal{T}(n)$ labeled by [0, n + 1] with root n + 1 is given by the formula

$$P(\mathcal{T}(n) = \mathbf{t}) = p_{n+1}^{|\mathbf{t}_{n+1}|-1} \prod_{i=0}^{n} p_i^{|\mathbf{t}_i|}$$
(3)

for all trees t labeled by [0, n + 1] with root n + 1, where $|\mathbf{t}_i|$ is the number of children of vertex *i* in the tree t. See [14, 15, 16, 17] for various constructions of $\mathcal{T}(n)$ and related results. For a vertex *x* of a rooted tree t let $V_x(\mathbf{t})$ be the set of non-root vertices of the *fringe subtree of* t rooted at *x*, that is the set of all vertices *v* of t in such that there is a path from *x* to *v* in t directed away from the root of t. See Aldous [3] for background and further references to fringe subtrees.

Theorem 1 [15]. For a random tree $\mathcal{T}(n)$ with distribution (3) on trees labeled by [0, n+1] with root n + 1, the random set $V(n) := V_0(\mathcal{T}(n))$ of non-root vertices of the fringe subtree of $\mathcal{T}(n)$ rooted at 0 has the Hurwitz distribution (2).

The purpose of this paper is to present the following result regarding the distribution of the size of such a random set V(n), in a limiting regime as $n \to \infty$ and the probability distribution p on [0, n + 1] is allowed to vary as a function of n.

Theorem 2 Let V(n) be a random subset of [n] with the Hurwitz distribution (2), for $(p_i := p_{i,n}, 0 \le i \le n+1)$ a probability distribution on [0, n+1] which varies with n. Then there exist limits

$$q_a = \lim_{n \to \infty} P(|V(n)| = a)$$
 for all $a = 0, 1, 2, \dots$ (4)

if and only if there exists a limit

$$\lambda = \lim_{n \to \infty} n p_{0,n} \in [0,\infty]$$

and one of the following conditions obtains: (i) $\lambda = 0$, in which case $(q_a) = (1, 0, 0, ...)$ (ii) $\lambda = \infty$, in which case $(q_a) = (0, 0, 0, ...)$ (iii) $\lambda \in (0, \infty)$ and the sequence of probability measures F_n on $[0, \infty)$ defined by

$$F_n[0,t] := n^{-1} \sum_{i=1}^n \mathbb{1}(np_{i,n} \le t)$$

has a weak limit F, a probability distribution on $[0,\infty)$ with mean at most 1; then

$$q_a = \frac{\lambda}{a!} \int_0^\infty (\lambda + t)^{a-1} e^{-(\lambda+t)} F^{a*}(dt) \qquad (a = 0, 1, 2, \ldots)$$
(5)

where F^{a*} is the a-fold convolution of F (in particular F^{0*} is a unit mass at 0 so that $q_0 = e^{-\lambda}$), and $\sum_a q_a = 1$.

In the particular case when $p_{i,n} = \theta/n$ for all $i \in [n]$ and some $0 \le \theta \le 1$ the distribution of |V(n)| is the quasi-binomial distribution studied by Consul [5, 6]. In this case $F_n = \delta_{\theta}$, a unit mass at θ , for every n. The limit distribution (q_a) in (5) for $F = \delta_{\theta}$ is the generalized Poisson distribution with parameters λ and θ :

$$q_a = \frac{\lambda}{a!} (\lambda + a\theta)^{a-1} e^{-(\lambda + a\theta)} \qquad (a = 0, 1, \ldots).$$
(6)

So Theorem 2 is a generalization of Consul's description of this generalized Poisson distribution as a limit of quasi-binomial distributions. Let P_{λ} denote the usual Poisson distribution with parameter λ , which is the case $\theta = 0$ of (6). In the setting of Theorem 1, as $n \to \infty$ and $np_{0,n}$ remains bounded, it can be shown using results of [15] that the delabeled fringe sub-tree of $\mathcal{T}(n)$ rooted at 0 is well approximated by a modified Galton-Watson branching process in which the root individual has offspring distribution P_{λ} for $\lambda = np_{0,n}$, and all following individuals have offspring distribution that is the mixed Poisson distribution $\int_0^{\infty} P_{\lambda} F_n(d\lambda)$. Details of this approximation will not be given here, but for similar approximations of combinatorially defined random trees by Galton-Watson trees, and related constructions, see [9, 2, 3, 1]. The Galton-Watson approximation to the delabeled fringe sub-tree of $\mathcal{T}(n)$ rooted at 0 explains why the limit distribution (q_a) in (5) has the the following interpretation as a Lagrangian distribution [13, 20]: **Theorem 3** For each probability distribution F on $[0, \infty)$, and each $\lambda \geq 0$, the sequence (q_a) defined by (5) is the distribution of the total progeny $\sum_{i=0}^{\infty} Z_i$ in a Galton-Watson branching process (Z_0, Z_1, \ldots) where the initial number of individuals Z_0 has the Poisson (λ) distribution P_{λ} , and the offspring distribution is the mixture of Poisson distributions $\int_0^{\infty} P_{\lambda} F(d\lambda)$.

Note that

$$\sum_{a} q_a = P(\sum_{i} Z_i < \infty) \tag{7}$$

which equals 1 iff F has mean at most 1, by standard theory of branching processes.

2 Proofs

The interpretation of (q_a) given in Theorem 3 is needed for the proofs of some parts of Theorem 2, so Theorem 3 will be proved first.

Proof of Theorem 3. According to a formula of Otter [12] and Dwass [7], which is reviewed in [17], for a Galton-Watson process (Z_0, Z_1, \ldots) with arbitrary offspring distribution, for every $k = 1, 2, \ldots$ and $a = 1, 2, \ldots$

$$P(\sum_{i} Z_{i} = a \mid Z_{0} = k) = \frac{k}{a} P(S_{a} = a - k)$$

where S_a is the sum of *a* independent copies of the generic offspring variable. If the offspring distribution is $\int_0^\infty P_\lambda F(d\lambda)$, then S_1 has the same distribution as $N(T_1)$ where $(N(t), t \ge 0)$ is a homogeneous Poisson process independent of T_1 with distribution F. Hence S_a has the same distribution as $N(T_a)$ for T_a the sum of *a* independent copies of T_1 , so $P(T_a \in dt) = F^{a*}(dt)$. So for $1 \le k \le a$

$$P(S_a = a - k) = P(N(T_a) = a - k) = \int_0^\infty e^{-t} \frac{t^{a-k}}{(a-k)!} F^{a*}(dt).$$

It follows that for Z_0 with Poisson(λ) distribution and a = 1, 2, ...

$$P(\sum_{i} Z_{i} = a) = \int_{0}^{\infty} f(a, \lambda, t) F^{a*}(dt)$$

where

$$f(a,\lambda,t) := \sum_{k=1}^{a} e^{-\lambda} \frac{\lambda^{k}}{k!} \frac{k}{a} e^{-t} \frac{t^{a-k}}{(a-k)!} = \frac{\lambda}{a!} (\lambda+t)^{a-1} e^{-(\lambda+t)},$$

and the conclusion of Theorem 3 follows.

Proof of Theorem 2. Let V(n) be a random subset of [n] with the Hurwitz distribution (2) induced by $(p_i = p_{i,n}, 0 \le i \le n+1)$. So for $a \in [0, n]$

$$P(|V(n)| = a) = \sum_{A \subseteq [n]:|A|=a} p_0 (p_0 + p_A)^{a-1} (p_{n+1} + p_{\bar{A}})^{n-a}.$$
(8)

In particular

$$P(|V(n)| = 0) = (1 - p_{0,n})^n \to e^{-\lambda}$$

as $n \to \infty$ iff $np_{0,n} \to \lambda$ for some $\lambda \in [0, \infty]$. Cases (i), (ii) and (iii) now arise according to the value of λ . Consider next

$$P(|V(n)| = 1) = \sum_{i=1}^{n} p_{0,n} (1 - p_{0,n} - p_{i,n})^{n-1} = n p_{0,n} \int_{0}^{\infty} \left(1 - p_{0,n} - \frac{x}{n} \right)^{n-1} F_{n}(dx).$$

In case (iii) with $np_{0,n} \to \lambda$ it follows easily from this expression that

$$P(|V(n)| = 1) \sim \lambda \int_0^\infty \left(1 - \frac{\lambda + x}{n}\right)^{n-1} F_n(dx) \to \lambda \int_0^\infty e^{-\lambda - x} F(dx)$$

if F_n converges weakly to F. A similar approximation of multiple sums by multiple integrals shows that in the same limit regime as $n \to \infty$

$$P(|V(n)| = a) \to \frac{\lambda}{a!} \int_0^\infty \cdots \int_0^\infty \left(\lambda + \sum_{i=1}^a x_i\right)^{a-1} \exp\left(-\lambda - \sum_{i=1}^a x_i\right) F(dx_1) \cdots F(dx_a)$$

for every $a = 1, 2, \ldots$, which yields (5). Since F_n has mean at most 1 for every n, so does any weak limit F. It then follows from the sentence after Theorem 3 that $\sum_a q_a = 1$. Conversely, if a limit q_a of P(|V(n)| = a) exists for all a, then the limit λ of $np_{0,n}$ exists by consideration of a = 0. In case (iii), the sequence F_n is tight because of the upper bound 1 on the mean. By passing to subsequences, the limit (q_a) must be of the form (5) for some subsequential weak limit F. To see that this F is in fact the weak limit of the whole sequence F_n it is enough to show that (q_a) determines F uniquely, and this follows easily from Theorem 3.

References

 D. Aldous. Tree-valued Markov chains and Poisson-Galton-Watson distributions. In D. Aldous and J. Propp, editors, *Microsurveys in Discrete Probability*, number 41 in DIMACS Ser. Discrete Math. Theoret. Comp. Sci, pages 1-20, 1998.

- [2] D.J. Aldous. A random tree model associated with random graphs. Random Structures Algorithms, 1:383-402, 1990.
- [3] D.J. Aldous. Asymptotic fringe distributions for general families of random trees. Ann. Appl. Probab., 1:228-266, 1991.
- [4] A. Cayley. A theorem on trees. Quarterly Journal of Pure and Applied Mathematics, 23:376-378, 1889. (Also in The Collected Mathematical Papers of Arthur Cayley. Vol XIII, 26-28, Cambridge University Press, 1897).
- [5] P.C. Consul. A simple urn model dependent upon a predetermined strategy. Sankhya Ser. B, 36:391-399, 1974.
- [6] P.C. Consul. Generalized Poisson Distributions. Dekker, 1989.
- [7] M. Dwass. The total progeny in a branching process. J. Appl. Probab., 6:682-686, 1969.
- [8] J. Françon. Preuves combinatoires des identités d'Abel. Discrete Mathematics, 8:331-343, 1974.
- [9] G. R. Grimmett. Random labelled trees and their branching networks. J. Austral. Math. Soc. (Ser. A), 30:229-237, 1980.
- [10] A. Hurwitz. Uber Abel's Verallgemeinerung der binomischen Formel. Acta Math., 26:199-203, 1902.
- [11] D.E. Knuth. Discussion on Mr. Riordan's paper. In R.C. Bose and T.A. Dowling, editors, *Combinatorial Mathematics and its Applications*, pages 71–91. Univ. of North Carolina Press, Chapel Hill, 1969.
- [12] R. Otter. The multiplicative process. Ann. Math. Statist., 20:206–224, 1949.
- [13] A. G. Pakes and T. P. Speed. Lagrange distributions and their limit theorems. SIAM Journal on Applied Mathematics, 32:745-754, 1977.
- [14] J. Pitman. Coalescent random forests. Technical Report 457, Dept. Statistics, U.C. Berkeley, 1996. To appear in J. Comb. Theory A. Available via http://www.stat.berkeley.edu/users/pitman.

- [15] J. Pitman. Abel-Cayley-Hurwitz multinomial expansions associated with random mappings, forests and subsets. Technical Report 498, Dept. Statistics, U.C. Berkeley, 1997. Available via http://www.stat.berkeley.edu/users/pitman.
- [16] J. Pitman. The multinomial distribution on rooted labeled forests. Technical Report 499, Dept. Statistics, U.C. Berkeley, 1997. Available via http://www.stat.berkeley.edu/users/pitman.
- [17] J. Pitman. Enumerations of trees and forests related to branching processes and random walks. In D. Aldous and J. Propp, editors, *Microsurveys in Discrete Probability*, number 41 in DIMACS Ser. Discrete Math. Theoret. Comp. Sci, pages 163–180, Providence RI, 1998. Amer. Math. Soc.
- [18] A. Rényi. On the enumeration of trees. In R. Guy, H. Hanani, N. Sauer, and J. Schonheim, editors, *Combinatorial Structures and their Applications*, pages 355– 360. Gordon and Breach, New York, 1970.
- [19] L.W. Shapiro. Voting blocks, reluctant functions, and a formula of Hurwitz. Discrete Mathematics, 87:319-322, 1991.
- [20] M. Sibuya, N. Miyawaki, and U. Sumita. Aspects of Lagrangian probability distributions. Studies in Applied Probability. Essays in Honour of Lajos Takács (J. Appl. Probab.), 31A:185-197, 1994.
- [21] V. Strehl. Identities of the Rothe-Abel-Schläfli-Hurwitz-type. Discrete Math., 99:321-340, 1992.